Causal Inference via Quantifying Influences

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Quantifying Causal Strength



Causal Inference via Quantifying Influences

Confounder Detection



Quantifying Causal Strength



A Crash Course in Causality





Defining causality is difficult. Here are some attempts:

Counterfactuals

X causes Y if:

Interventions

Observations

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- X causes Y if:
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• $Var[Y_t | X_t, Y_{1:t-L}] < Var[Y_t | Y_{1:t-L}]$ [Granger 1969, Wiener 1956]



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- If we intervene on X, then Y changes [Pearl 2009, among others]



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Consider Simpson's paradox



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Consider Simpson's paradox

do(X = x) decreases Y



- Intervention is stronger than any purely statistical concept
- The notation for an intervention is do(X = x)
- Statistically, X is positively correlated with Y; but the intervention









These imply joint probabilities + data-generating factorization



- To model Pearl's causality, we use *Bayesian networks*
- These imply joint probabilities + data-generating factorization
 - $p(x, y, z) = p(y | x, z) \times p(x | z) \times p(z)$



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functionally

By reparameterizing, we can always rewrite the Bayesian network





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 $X := f(\operatorname{Pa}(X), \varepsilon)$





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"X is a function of its parents (causes) and a noise term"

By reparameterizing, we can always rewrite the Bayesian network

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Causal discovery is ill-posed for observational data



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Non-parametrically, no reason to favor the "true" factorization





Two (mutually-inclusive) options:



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Two (mutually-inclusive) options: Intervene on the system

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Two (mutually-inclusive) options: Intervene on the system

Make assumptions about the data-generating process

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Example assumptions:





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Lack of causations => statistical independence (Causal Markov)





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> All of these assumptions break symmetry



An Outline of Today

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Using these notions from causal inference, we'll talk about How to quantify causal strengths

- observational data
- Some concluding thoughts

• How to find "hidden common causes" using causal strength Using causal strength to discovery causal relationships from

Quantifying Causal Strength





Practically, we don't only care that a causal relation exists



Practically, we don't only care that a causal relation exists

The strength of a relationship is also important



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• Fraction of the variance of X_i which is controlled by X_i [ANOVA]



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For example:

- distribution [Information Flow]

• Fraction of the variance of X_i which is controlled by X_i [ANOVA] Kullback-Leibler divergence of marginal vs. conditional



Some interventional attempts:

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- Kullback-Leibler divergence of interventional distributions with/without specified edge [Janzing et al., AOS 2013]
- Differentiating the expected value under $do(X_i = x_i)$:



 $\frac{\partial}{\partial x_i} \mathbb{E}[X_j | \operatorname{do}(X_i = x_i)] \text{ [Average Causal Effect]}$



The linear case is instructive for us





 $X_j = \sum \beta_{i \to j} X_i + \varepsilon$ $X_i \in \operatorname{Pa}(X_j)$

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$$X_{j} = \sum_{X_{i} \in \operatorname{Pa}(X_{j})} \beta_{i \to j} X_{i} + \varepsilon$$

Then, it's intuitive to link $\beta_{i \rightarrow j}$ to "causal strength"

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This aligns with the average causal effect, is similar to ANOVA, and more



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How do we move to nonlinear case?



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A widely used measure is the average causal effect (ACE): $\mathsf{ACE}_{X_i \to X_j} \triangleq \frac{\partial}{\partial x_i} \mathbb{E}[X_j | \operatorname{do}(X_i = x_i)]$

SPL 2022]:





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Our alternative is the differential causal effect (DCE) [Butler et al.,

$$X_{j} = f(X_{1}, \dots, X_{N}, \varepsilon)$$
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SPL 2022]:



By differentiating under the integral sign, the ACE is the average DCE



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For strength $S_{X_i \to X_j}$, it's desirable that



The Average Causal Effect Has Issues





$\mathcal{S}_{X_i \to X_j} \neq 0 \iff X_i \to X_j$
The Average Causal Effect Has Issues



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The ACE fails this test!

This is the case whenever the DCE is zero-mean, i.e. $\mathbb{E}\left[\mathrm{DCE}_{X_i \to X_j}\right] = 0$



$$X_i \to X_j \neq 0 \iff X_i \to X_j$$



Wikimedia Commons, CC-BY-SA-4.0



How we estimate $DCE_{X_i \rightarrow X_j}$ depends on how we estimate f_j

Wikimedia Commons, CC-BY-SA-4.0



"ForwardAD.png" by MaxEmanuel, Wikimedia Commons, CC-BY-SA-4.0

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For many estimators, this is available in closed-form



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How we estimate $DCE_{X_i \rightarrow X_i}$ depends on how we estimate f_i

- For many estimators, this is available in closed-form
- Even more generally, automatic differentiation makes this easy

Confounder Detection

Y. Liu, C, Cui, D. Waxman, K. Butler, and P. M. Djurić, "Detecting confounders in multivariate time series using strength of causation," Proceedings of the 31st European Signal Processing Conference, Helsinki, Finland, 2023.





Definition: if there exists a variable Z causing at least two other variables, it is known as a confounder





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One of the most common assumptions is causal sufficiency, i.e. there are no latent confounders







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ordinary regression



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- This disallows causes from the future affecting the present
- If sampled with sufficiently high rate, we can also disallow
- The assumption of no instantaneous causes turns discovery into





Structure Learning



Structure Learning

Finds a modified *causal graph* with possible \bullet confoundedness indicated



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 - Information-theoretic techniques (e.g., [Kaltenpoth & Vreeken, SDM 2019])
- Our work: extension to time series using *latent variable models (LVMs) and differential causal effect (DCE)*





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- It is possible to infer what we can't see (up to diffeomorphism)

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Idea:

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Learn an LVM

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Idea: 1. Learn a latent variable model (LV)

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Idea:

- **1.** Learn a latent variable model (LV) **2.** Perform inference of latent time series Z_t

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Idea:

- **1.** Learn a latent variable model (LV) **2.** Perform inference of latent time series Z_t
- **3.** Test the DCE of z_{t-i} to y_t

- It is possible to infer what we can't see (up to diffeomorphism)




in multivariate time series

The exact LVM is not so important, but we desire online learning





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We use deep Gaussian process state-space models [Liu et al., TSP] 2023]







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These write time-series auto regressively with Gaussian processes (GPs)

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processes (GPs)

and the GP parameters

- The exact LVM is not so important, but we desire online learning
- We use deep Gaussian process state-space models [Liu et al., TSP]
- These write time-series auto regressively with Gaussian

Using a specific GP approximation, they filter on the latent state









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 $\mathbf{Z}_t = f(\mathbf{Z}_t)$ $\mathbf{x}_t = h(\mathbf{z}_t)$ $y_t = g(\mathbf{Z}_t)$

Then with unknown functions f, g, h and white Gaussian noise

$$t - l_{zz}:t-1, \mathbf{X}_{t-l_{zx}}:t-1, y_{t-l_{zy}}:t-1) + \mathbf{u}_{t},$$

$$t - l_{xz}:t-1, \mathbf{X}_{t-l_{xx}}:t-1, y_{t-l_{xy}}:t-1) + \mathbf{v}_{t},$$

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Any LVM that learns f, g, h and \mathbf{z}_t is good for us

Then with unknown functions f, g, h and white Gaussian noise

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2. Infer $DCE_{z_{k,t-i} \to y_t}$ for k = 1, ..., K and $i = 1, ..., l_{y_z} - 1$

Learn an LVM

Inference

Test DCE

In the inference step, two goals: 1. Infer z_t for t = 1, ..., T

In our case, $DCE_{z_{k,t-i} \rightarrow y_t}$ is available in a convenient closed form

2. Infer $DCE_{z_{k,t-i} \to y_t}$ for k = 1, ..., K and $i = 1, ..., l_{y_z} - 1$





or $DCE_{z_{k,t-i} \rightarrow y_t} \neq 0$ (influence)

For testing, the goal is to decide if $DCE_{z_{k,t-i} \rightarrow y_t} = 0$ (no influence)





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- For testing, the goal is to decide if $DCE_{z_{k,t-i} \rightarrow y_t} = 0$ (no influence)

1. Test if the p % credible interval of DCE $_{z_{k,t-i} \rightarrow y_t}$ contains 0 2. Test if $\Pr\left(DCE_{z_{k,t-i} \to y_t} \in (-\epsilon, \epsilon)\right)$ exceeds some threshold







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No...

coordinates:

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 ∂Z_{t-i}

No...

But for scalars, zeroness of the DCE is invariant to a change in coordinates:

Is the causal strength of a latent variable well-defined?

$$= \frac{\partial y_t}{\partial z'_{t-i}} \frac{\partial z'_{t-i}}{\partial z_{t-i}} = 0 \times \frac{\partial z'_{t-i}}{\partial z_{t-i}} = 0.$$



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$$= \frac{\partial y_t}{\partial z'_{t-i}} \frac{\partial z'_{t-i}}{\partial z_{t-i}} = 0 \times \frac{\partial z'_{t-i}}{\partial z_{t-i}} = 0.$$

Therefore, testing for zero-ness is well-defined

Our Method Can Detect Confounders In the Static Case







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Non-zero DCE for confounder





Our Method Can Detect Confounders In the Dynamic Case



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Our Method Can Detect Confounders In the Dynamic Case



Causal Discovery

D. Waxman, K. Butler, and P. M. Djurić "DAGMA-DCE: Interpretable, Non-Parametric Differentiable Causal Discovery" Submitted.

Discovering Causal Relationships Requires Assumptions



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causal graph



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causal graph

As before: assumptions, assumptions, and more assumptions



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As before: assumptions, assumptions, and more assumptions

Largely, two categories: 1. Constraint-based methods 2. Score-based methods



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Constraint-Based Methods Exploit

Independencies



C. Glymour, K. Zhang, & P. Spirtes. (2019). Review of causal discovery methods based on graphical models. *Frontiers in genetics*, *10*, 524.



The most historically popular methods are constraint-based

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causal relationship

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- The most historically popular methods are constraint-based
- Recall the "causal Markov" and "causal faithfulness assumptions" Together, statistical conditional independence if and only if encoded by a



causal relationship

This can falsify many causal graphs

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Inference (FCI)



- The most historically popular methods are constraint-based
- Recall the "causal Markov" and "causal faithfulness assumptions" Together, statistical conditional independence if and only if encoded by a
- Constraint-based methods test conditional independencies
- Famous examples include the PC Algorithm and Fast Causal



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Still, they can be a great tool

- But we'll work on score-based methods instead

MLE-type (or AIC/BIC-type) procedures are very nice

Surprisingly mild assumptions ensure identifiability

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- Gaussian noise

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- Gaussian noise

AIC/BIC

 Linear model with additive non-Gaussian noise (LiNGAM) • Three-times differentiable, strictly nonlinear with additive

Score-based methods search over DAGs to minimize the MLE/

- 1: 1
- 2: 3
- 3: 25
- 4: 543
- 5: 29281

• • •

14: 1.4×10³⁶

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The space of DAGs grows super-exponentially

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The space of DAGs grows super-exponentially

- 1: 1
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The space of DAGs grows super-exponentially

Let $\mathbf{A} =$ binary adjacency matrix of \mathcal{G}



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Let's look at k = 2:



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Searches over DAGs can then be constrained, continuous optimization



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Linear NOTEARS

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Use the linear coefficient mi Markovski s.t. tr

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in $(\exp(\mathbf{A} \odot \mathbf{A})) - d = 0$

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This is well-posed and gets very nice results

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$$\|\mathbf{X} - \mathbf{X}\mathbf{A}\|_{F}^{2} + \lambda \|\mathbf{A}\|_{1}$$

$$\int \left(\exp\left(\mathbf{A} \odot \mathbf{A}\right)\right) - d = 0$$

Nonlinear case [Zheng et al., AISTATS 2020] is similar

Idea: define $[\mathbf{A}]_{ij} = ||\partial_i f_j||_2$

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For parameterized m m1n θ s.t. tr

Nonlinear case [Zheng et al., AISTATS 2020] is similar

 $\|\partial_i f_j\|_2$

odel
$$\mathcal{M}_{\theta}$$

n $\|\mathbf{X} - f_{\theta}(\mathbf{X})\|_{2}^{2} + \lambda \|\mathbf{A}_{\theta}\|_{1}$
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In this case, M-matrices from econometrics

- Idea: define a class of matrices so that barrier methods work
- DAGMA [Bello et al., NeurIPS 2022] then gives a different constraint:
 - $\min_{\theta} \|\mathbf{X} f_{\theta}(\mathbf{X})\|_{F}^{2} + \lambda \|\mathbf{A}_{\theta}\|_{1}$ s.t. $-\log\left(\det\left(s\mathbb{I} \mathbf{A}_{\theta} \odot \mathbf{A}_{\theta}\right)\right) + d\log s = 0$

A_{θ} is Arbitrarily Misspecified



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DAGMA-MLP defines \mathbf{A}_{θ} using the L^2 norm

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but $\|\partial_i f_j\|_2 > \delta$.

DAGMA-MLP defines \mathbf{A}_{θ} using the L^2 norm

- Lemma: There exists an MLP with weight matrices $B^{(1)}, \ldots, B^{(M)}$ and sigmoidal activation such that $||B_1^{(1)}||_2 < \epsilon$

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compensate with very large edges in $B^{(2)}$

DAGMA-MLP defines \mathbf{A}_{θ} using the L^2 norm

- Lemma: There exists an MLP with weight matrices $B^{(1)}, \ldots, B^{(M)}$ and sigmoidal activation such that $||B_1^{(1)}||_2 < \epsilon$

Proof Idea: for each outgoing edge of $B^{(1)}$ which is small,



| -╂- | =

 $\|\partial_i f_j\| \equiv$

Define instead $\mathbf{A}_{\theta} \triangleq \|\partial_i f_j\|_{L_2(\mathbb{P}^X)}$



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We can take a Monte Carlo approximation





$$]_{ij} \approx \sqrt{\frac{1}{N} \sum_{n=1}^{N} \left(\partial_i f_j(x_n)\right)^2}$$



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We can take a Monte Carlo approximation



+ = $\|\partial_i f_j\| =$

This is the root-mean-square DCE

 $\left[\mathbf{A}_{\theta}\right]_{ij} \approx \sqrt{\frac{1}{N}\sum_{n=1}^{N} \left(\partial_{i}f_{j}(x_{n})\right)^{2}}$





Our optimization problem stays the same

$\min_{\theta} \|\mathbf{X} - f_{\theta}(\mathbf{X})\|_{2}^{2} + \lambda \|\mathbf{A}_{\theta}\|_{1}$ s.t. $-\log\left(\det\left(s\mathbb{I} - \mathbf{A}_{\theta} \odot \mathbf{A}_{\theta}\right)\right) + d\log s = 0$

DAGMA-DCE

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Notably, $\|\mathbf{A}_{\theta}\|_{1}$ is different!

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DAGMA-DCE

Our optimization prob mir s.t. $-\log\left(d\right)$

Notably, $\|\mathbf{A}_{\theta}\|_1$ is different!

Whenever the DCE is well-defined and easy to compute, terms in this problem are too

olem stays the same

$$\|\mathbf{X} - f_{\theta}(\mathbf{X})\|_{2}^{2} + \lambda \|\mathbf{A}_{\theta}\|_{1}$$

$$\det \left(s\mathbb{I} - \mathbf{A}_{\theta} \odot \mathbf{A}_{\theta}\right) + d\log s = 0$$

DAGMA-DCE Recovers Linear Strength

Data was generated with a linear SEM

Difference in Estimated





DAGMA-DCE Maintains Performance







Data was generated with additive Gaussian processes

... Even in Unfavorable Comparisons

identifiability







Data was generated with MLPs, made to ensure DAGMA

DAGMA-DCE Orders Variables Differently



Kendall's Tau	Spearman's Rho
0.40 ± 0.09	0.53 ± 0.11
0.55 ± 0.06	0.74 ± 0.07

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Both Kendall's τ and Spearman's ρ indicate different orderings



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thresholding



One of the ad-hoc components of NOTEARS+/DAGMA was

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DAGMA-DCE still thresholds, but the threshold is interpretable



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One of the ad-hoc components of NOTEARS+/DAGMA was

- DAGMA-DCE still thresholds, but the threshold is interpretable
- This allows the expert to decide what's a relevant effect

Concluding Remarks




Causal inference is important to our understanding of signals, systems, and their scientific context



Causal relationships have not only a direction, but a strength

Causal inference is important to our understanding of signals,



Strength is often thrown away

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Causal inference is important to our understanding of signals, systems, and their scientific context

Causal relationships have not only a direction, but a strength

Strength is often thrown away

By incorporating strength, we could Detect confounders in multivariate time series Increase interpretability in differentiable causal discovery









Interesting avenues with "hybrid" causal discovery methods



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- Lots of other places to use ML and causal strength in causality
- Interesting avenues with "hybrid" causal discovery methods
- Interpretability brings opportunities for "workflows"



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- Lots of other places to use ML and causal strength in causality
- Interesting avenues with "hybrid" causal discovery methods
- Interpretability brings opportunities for "workflows"
- Together, these empower decision-makers

Thank You!









Kurt Butler

Collaborators







Yuhao Liu

